



# The power of replications in difference tests

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## Result

**A few reps á day makes the low  
power go away!**

# Setting

- ◆ Panel size:  $n$
- ◆ Each assessor perform  $k$  difference(triangle) test
- ◆ The  $N=nk$  binomial test is OK (Meyners & Kunert, 1999)
- ◆ The power of this test is unknown!

# Outline

- ◆ Introduce the power of the usual binomial test with replications
- ◆ Present and compare different statistical models for replications
- ◆ Calculate and compare the power within these models
- ◆ Give limits for this power

# The power of the binomial test

- ◆ The probability of claiming a difference when in fact it is there:
  - $P(X=x_{\text{critical}})$ , where
  - $X$ =total number of correct answers
  - and assuming some "alternative situation":
  - some (at least one) on the individual  $p_i$ s are larger than  $1/3$ .

# Model types for the alternative

- ◆ Beta-binomial (Ennis & Bi, 1998)
- ◆ Generalized Linear Mixed Models (Brockhoff, 1997, Hunter et al., 2000)
- ◆ Binomial mixture models (Meyners & Kunert, 1999)

# The Beta-binomial model:

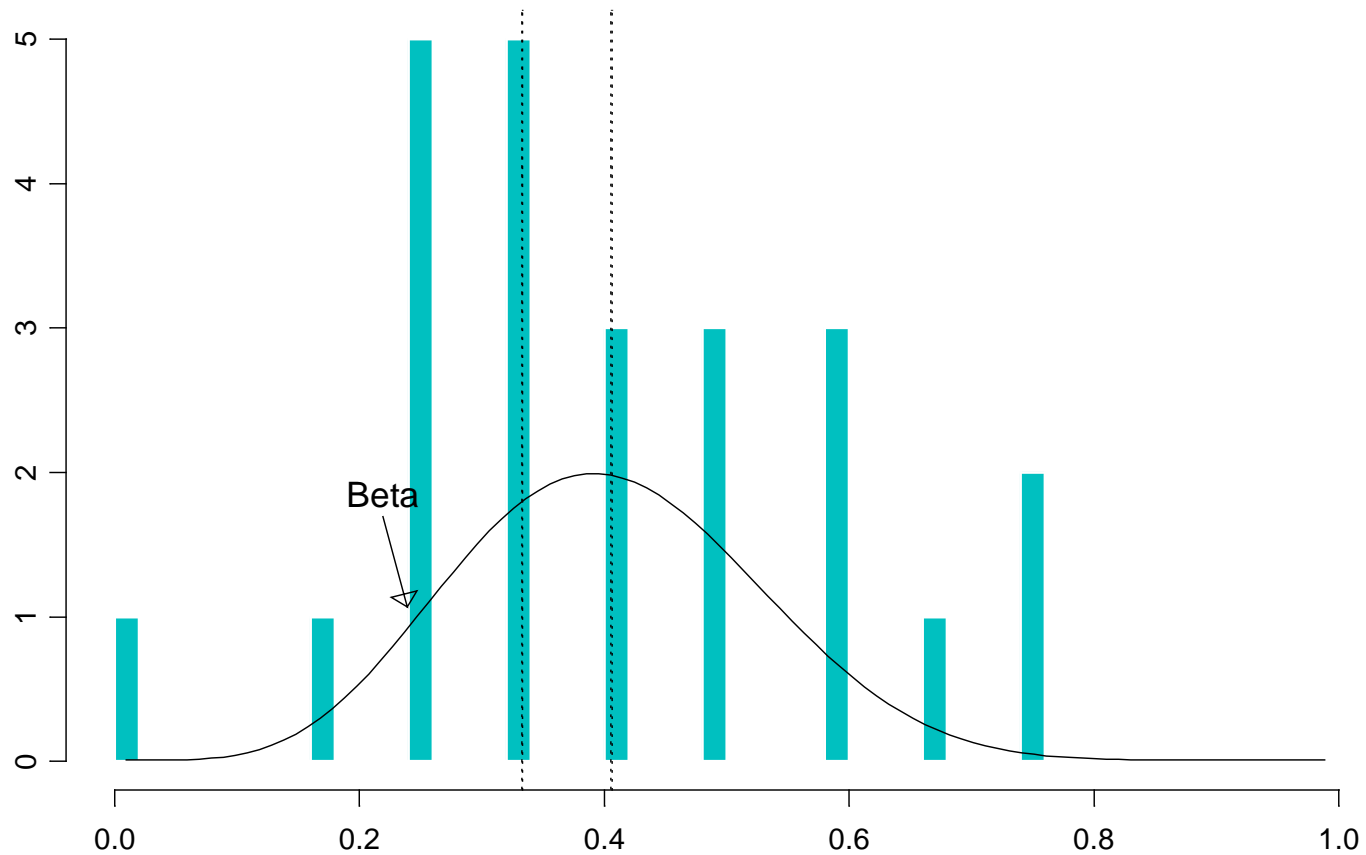
For each assessor:

$$X_i \sim \text{bin}(k, p_i)$$

The individual probabilities are randomly distributed:

$$p_i \sim \text{Beta}(\alpha, \beta)$$

Data:  $n=24$ ,  $k=12$



# The Generalized Linear Mixed Model (GLMM)

For each assessor:

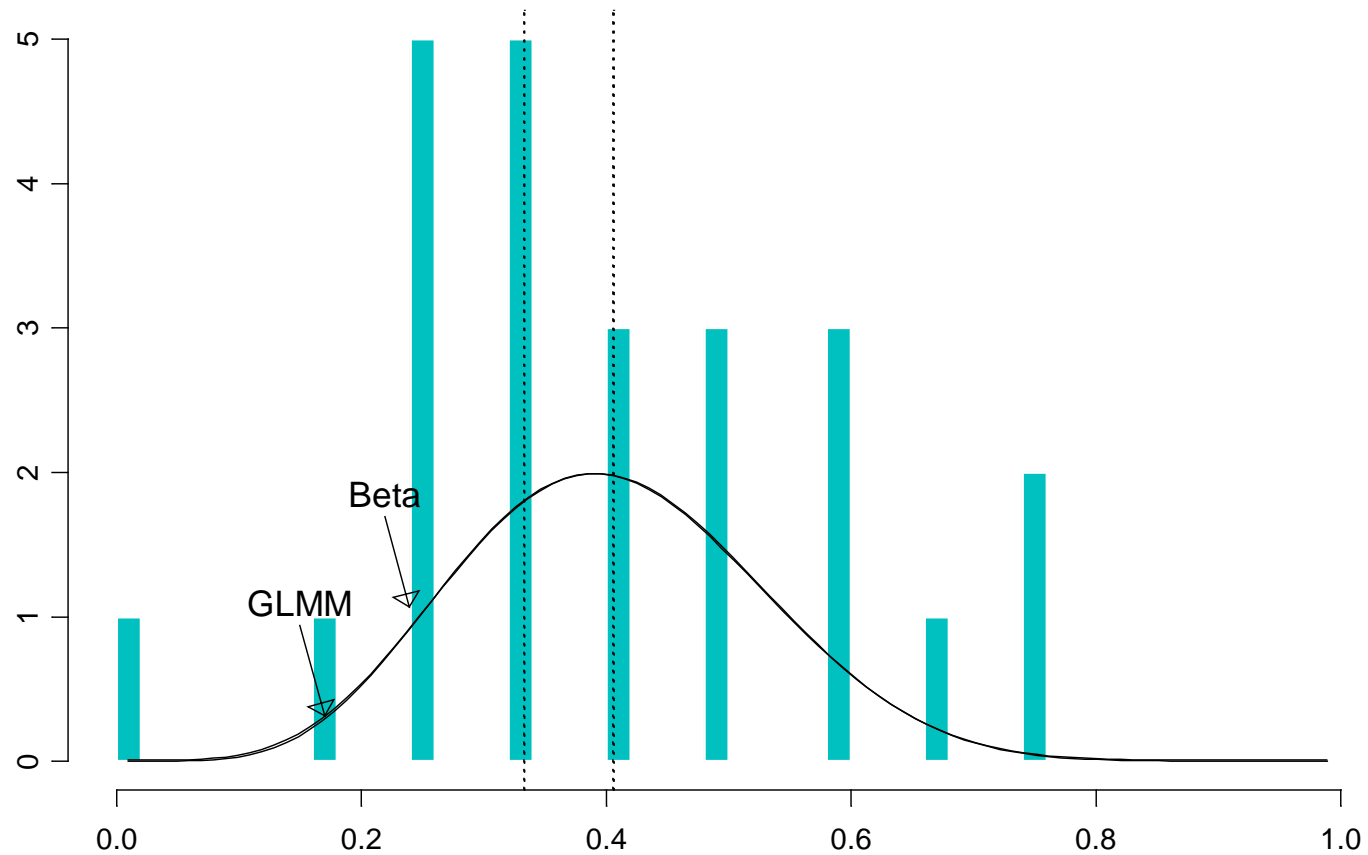
$$X_i \sim \text{bin}(k, p_i)$$

The individual probabilities are randomly distributed:

$$p_i = \Phi(d + Z_i)$$

$$Z_i \sim N(0, \sigma^2)$$

Data:  $n=24$ ,  $k=12$



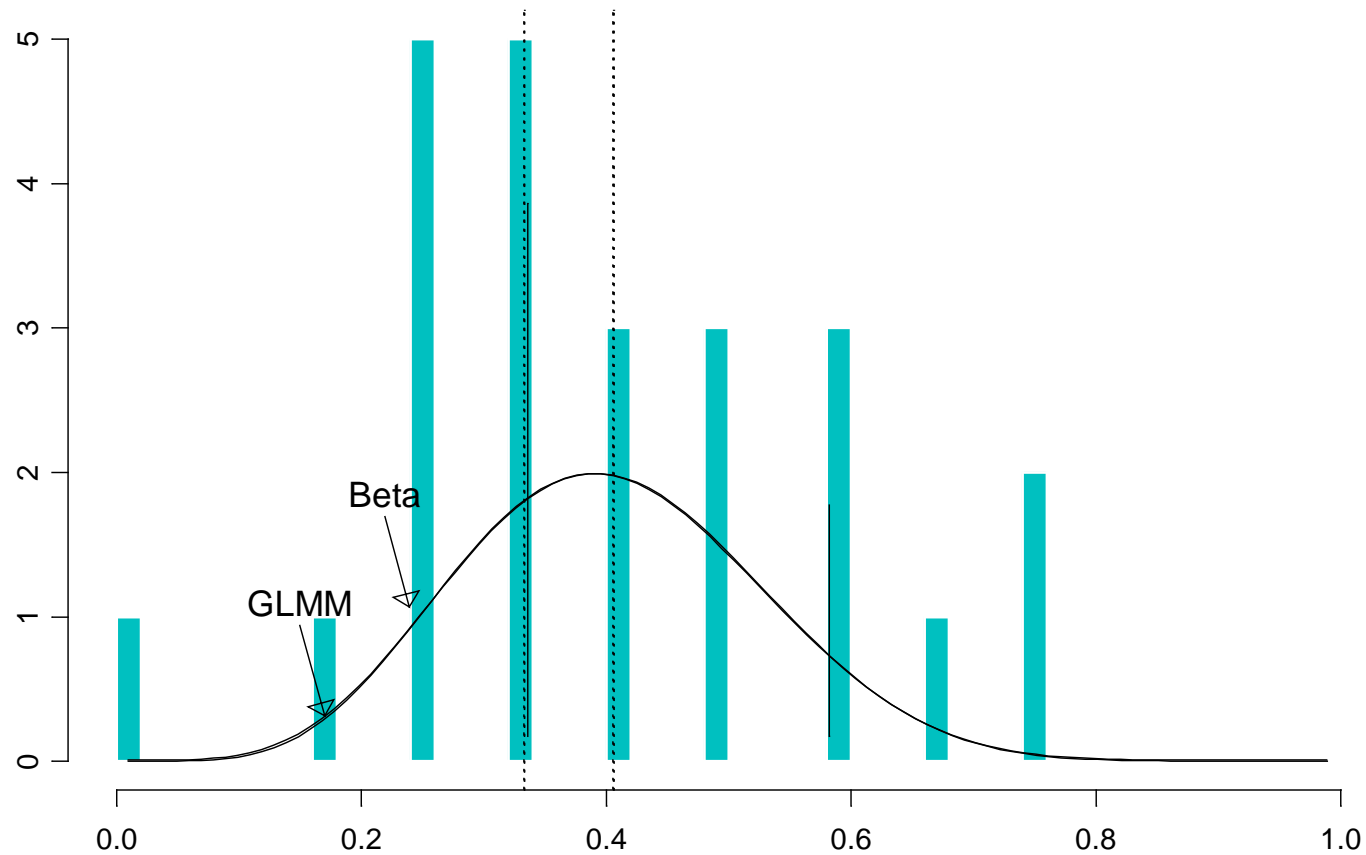
# The binomial mixture model

For each assessor:  $X_i \sim \text{bin}(k, p_i)$

The individual probabilities are randomly distributed:

$$p_i = \begin{cases} \frac{1}{3} & \text{with probability } 1-\pi \\ \frac{1}{3} + \left(1 - \frac{1}{3}\right)\pi & \text{with probability } \pi \end{cases}$$

Data:  $n=24$ ,  $k=12$



# The Corrected Beta-binomial model:

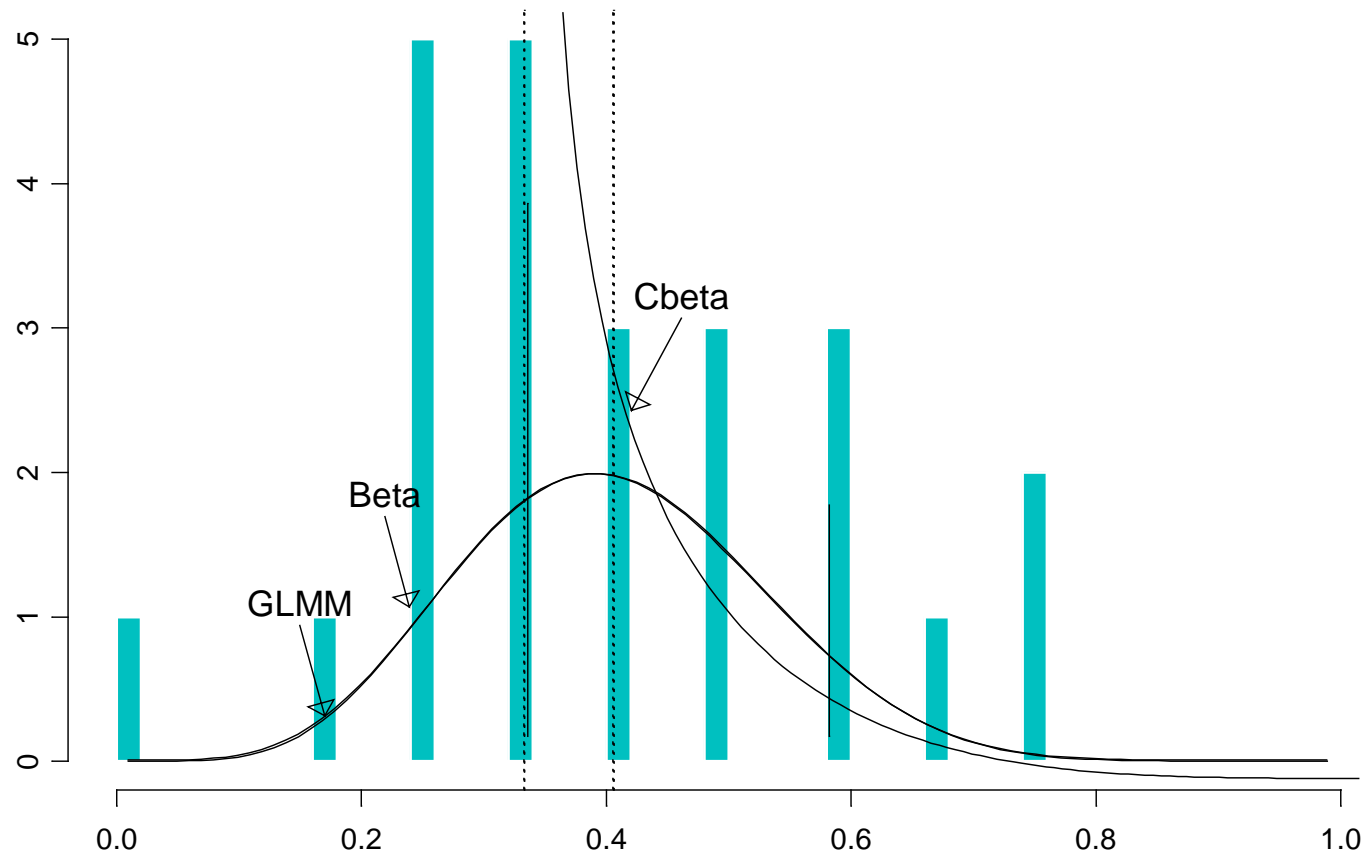
For each assessor:  $X_i \sim \text{bin}(k, p_i)$

The individual probabilities are randomly distributed:

$$p_i = \frac{1}{3} + \left(1 - \frac{1}{3}\right)\pi_i$$

$$\pi_i \sim \text{Beta}(\alpha, \beta)$$

Data:  $n=24$ ,  $k=12$



# The Corrected Generalized Linear Mixed Model (CGLMM)

For each assessor:

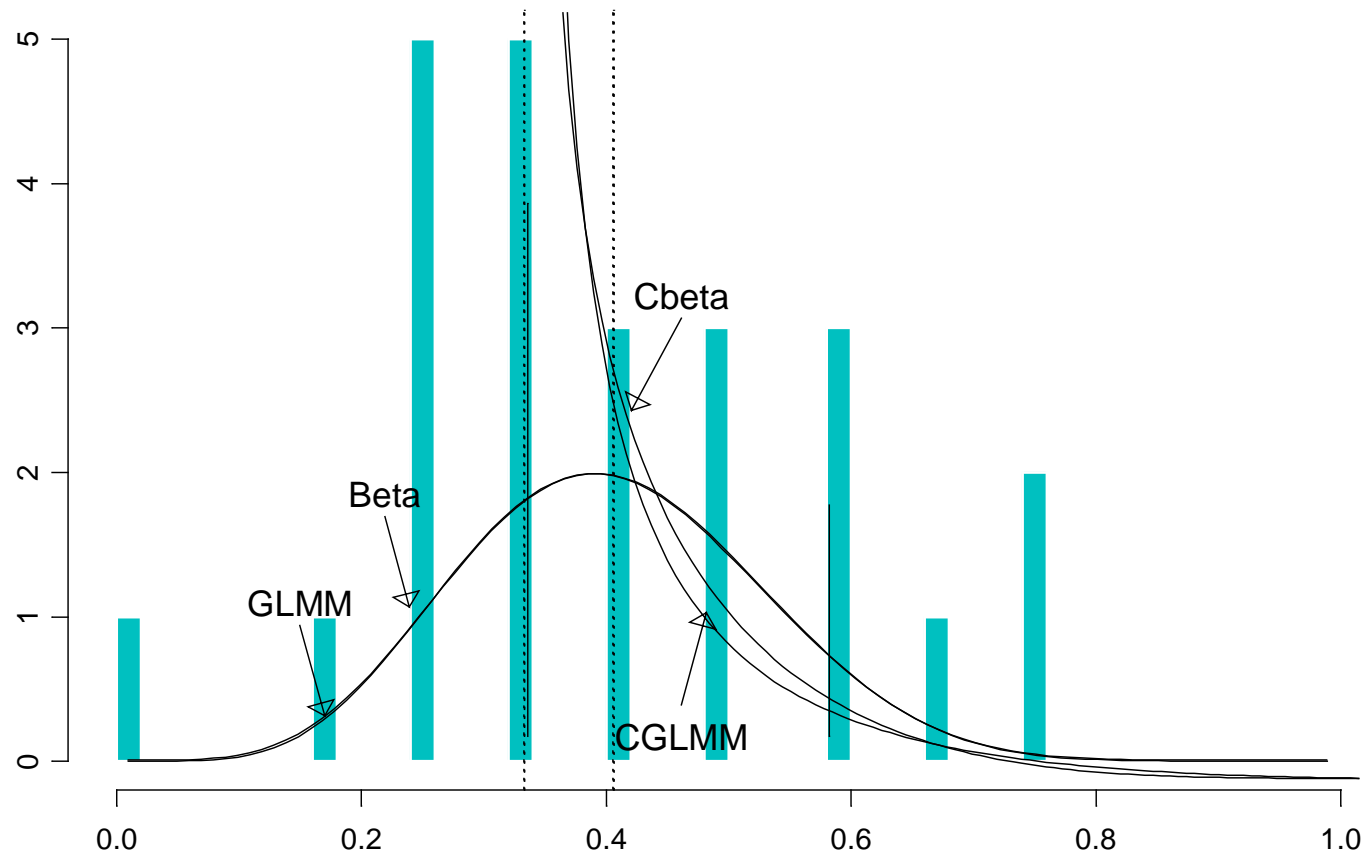
$$X_i \sim \text{bin}(k, p_i)$$

The individual probabilities are randomly distributed:

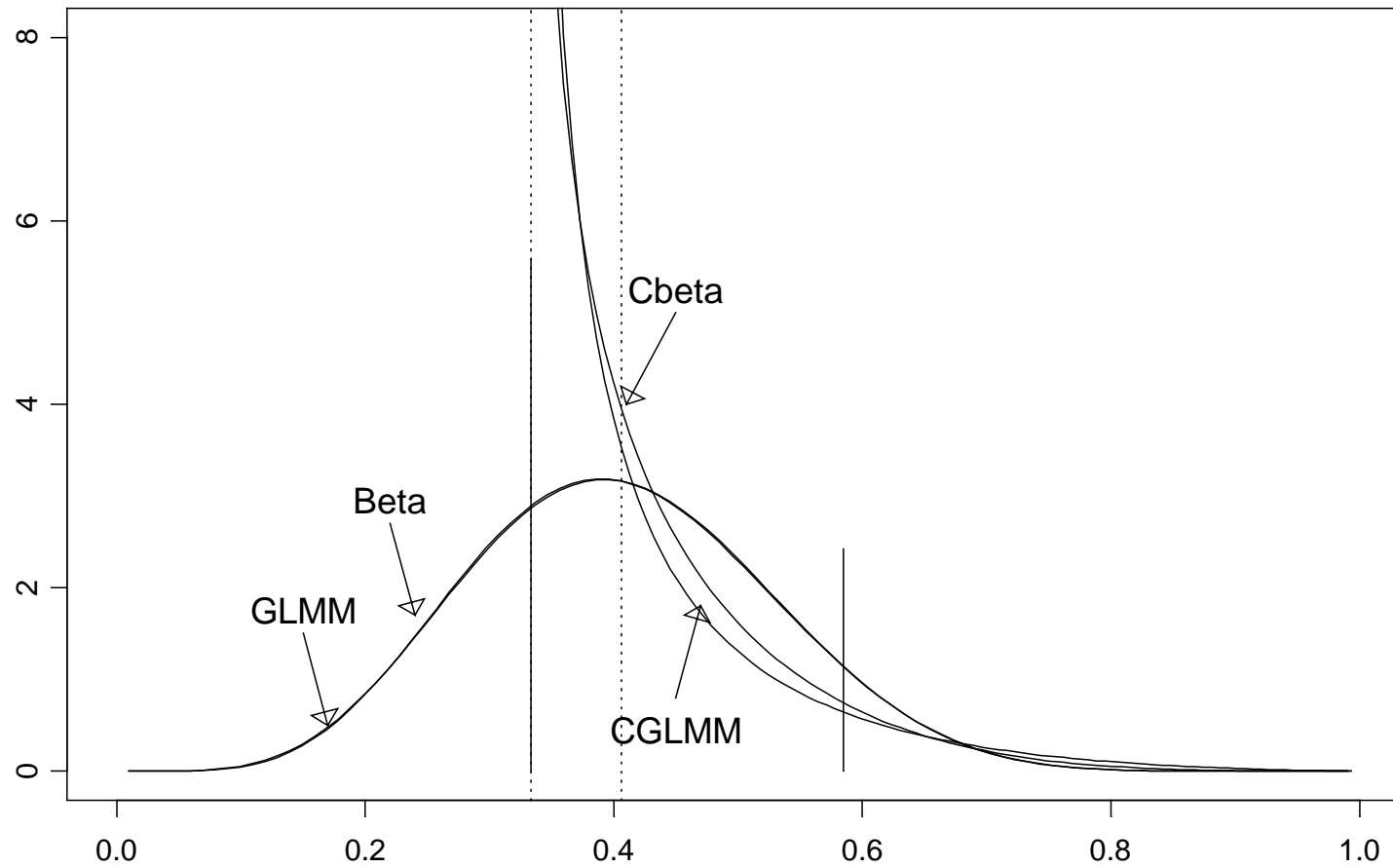
$$p_i = \frac{1}{3} + \left(1 - \frac{1}{3}\right)\Phi(d + Z_i)$$

$$Z_i \sim N(0, \sigma^2)$$

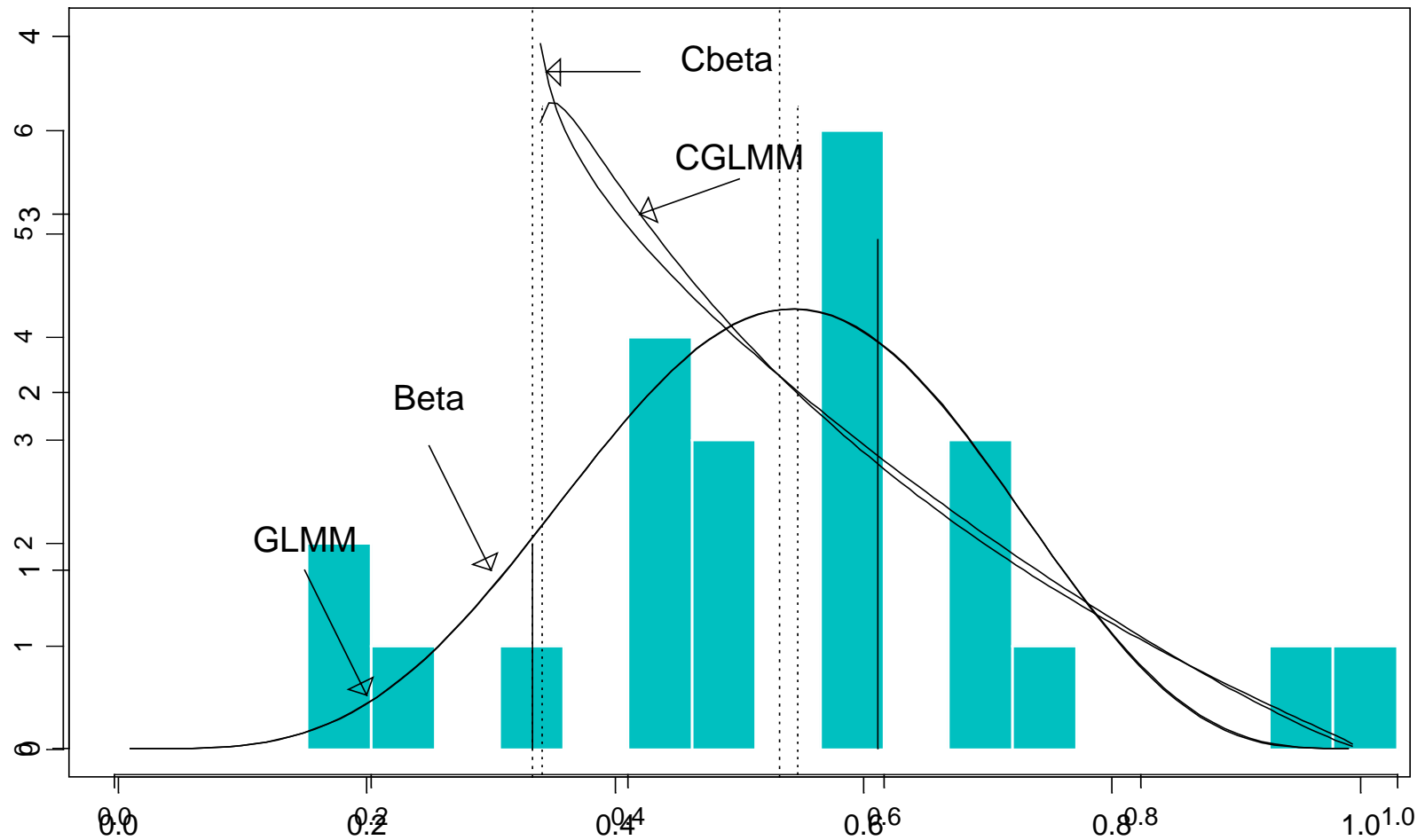
Data:  $n=24$ ,  $k=12$



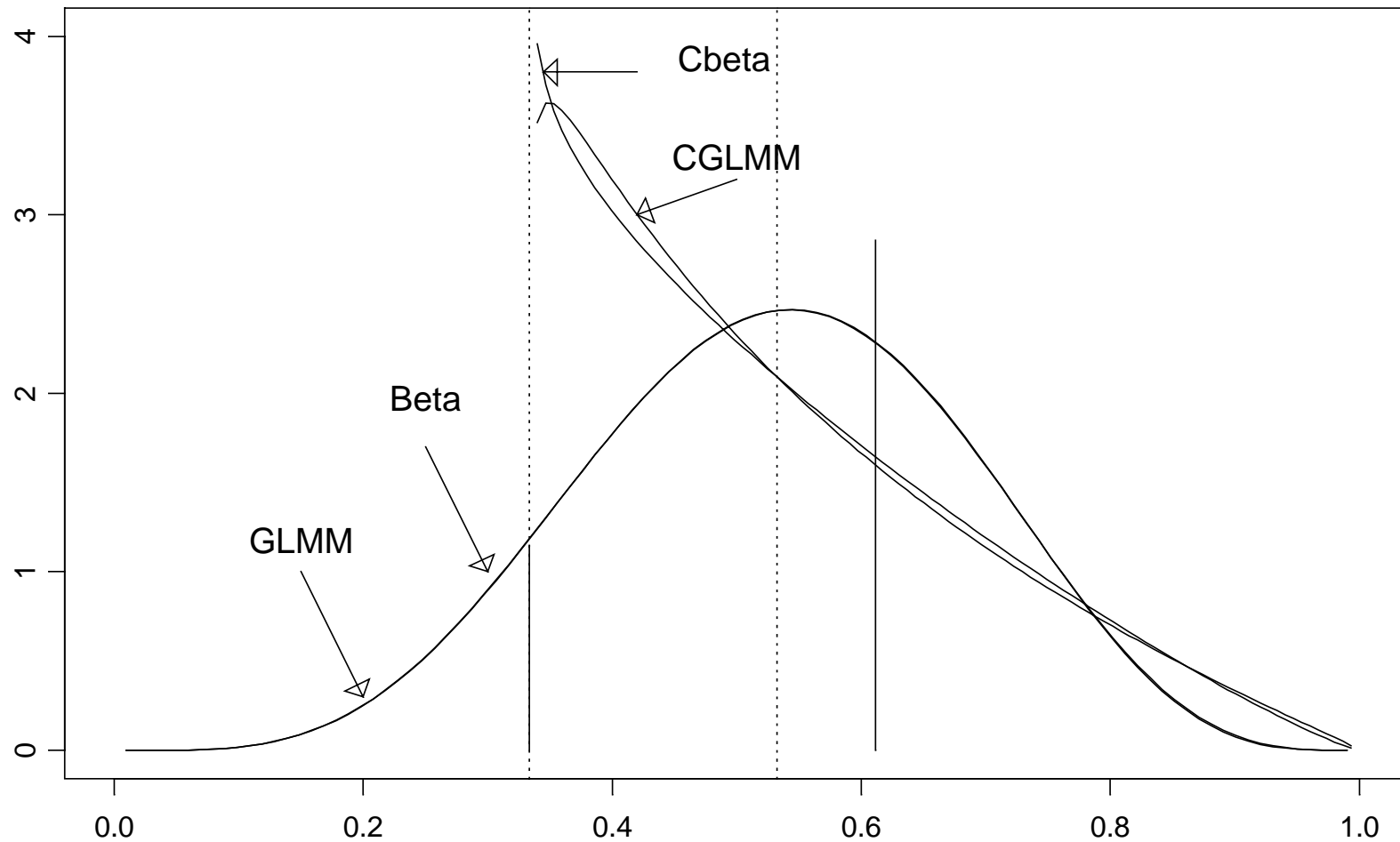
# Model comparison



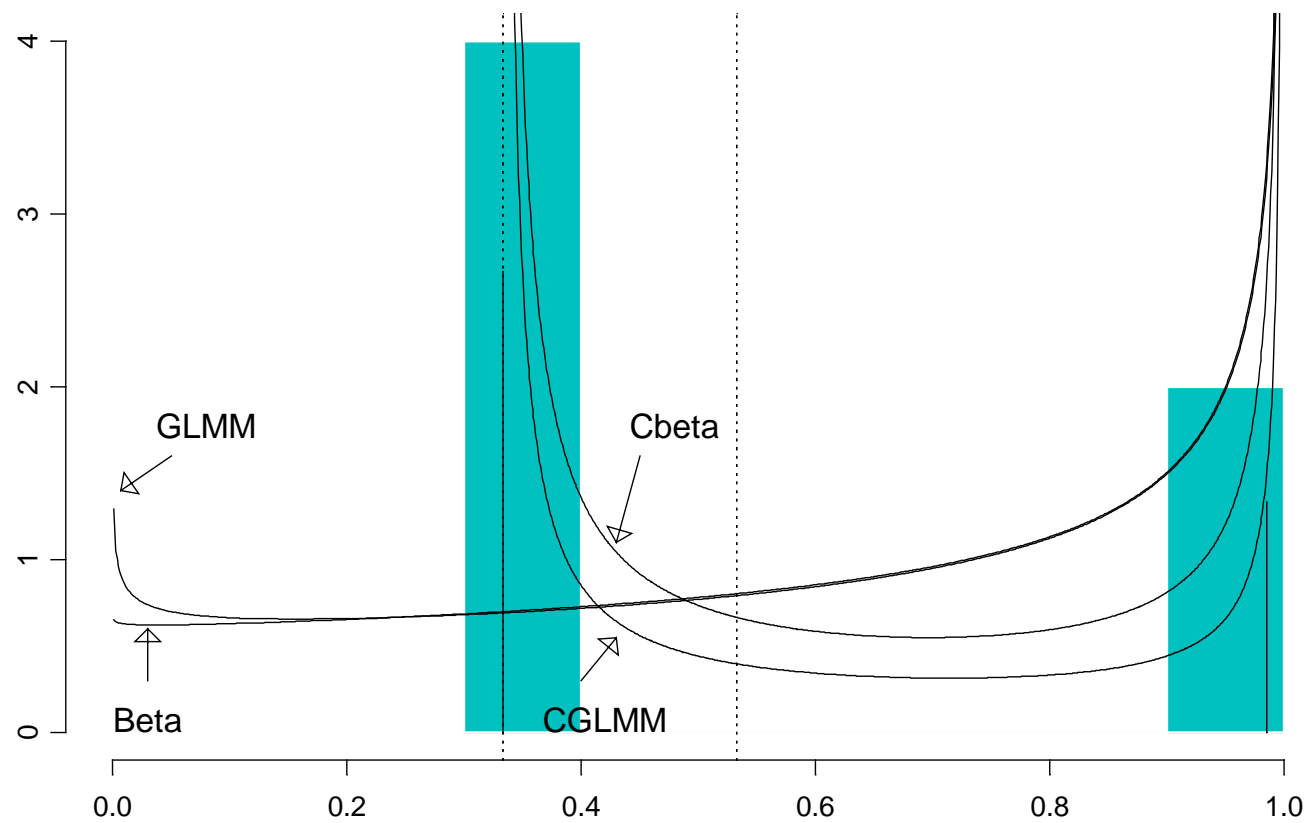
Data:  $n=23$ ,  $k=12$



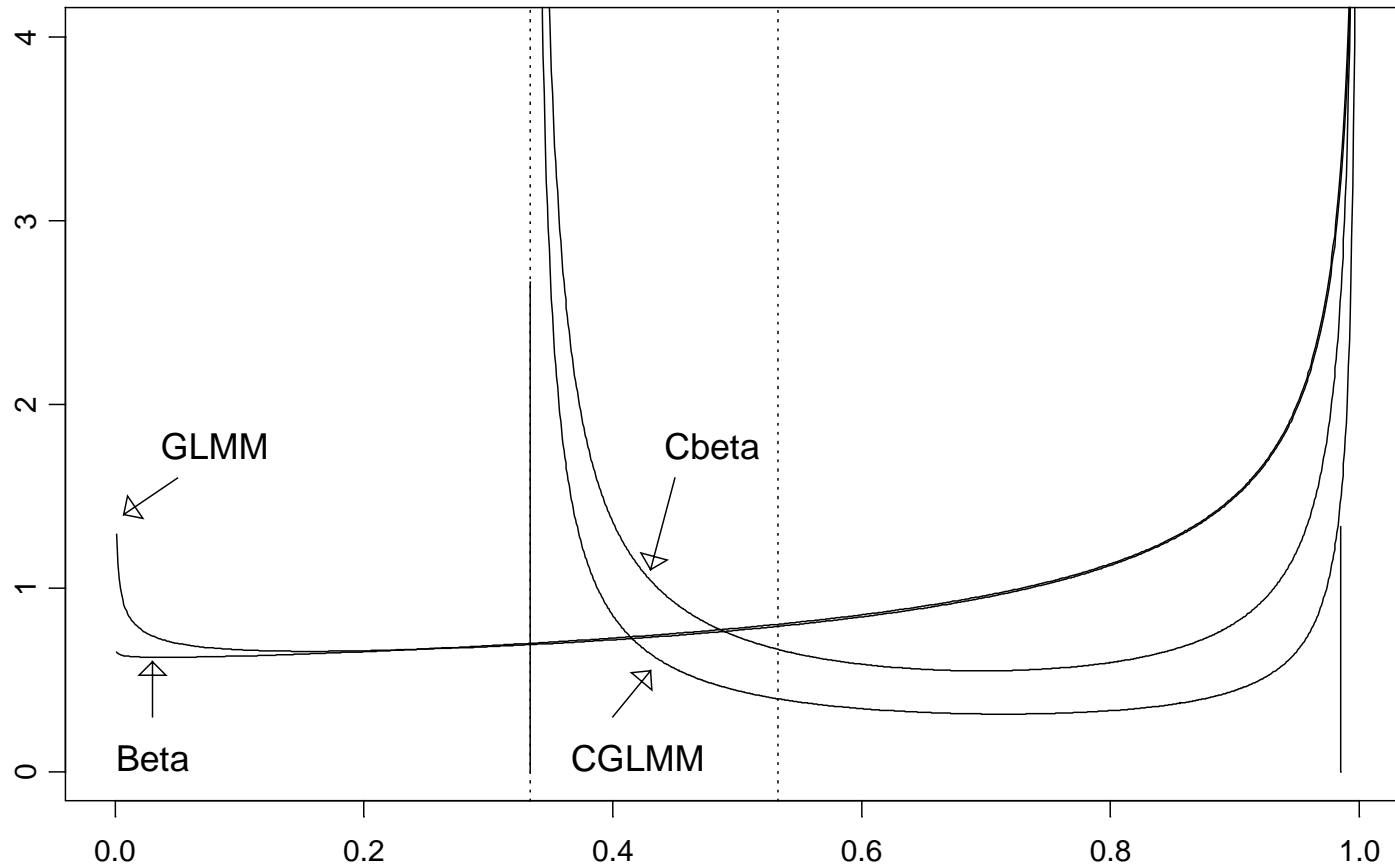
# Model comparison



Data:  $n=6$ ,  $k=100$



# Model comparison

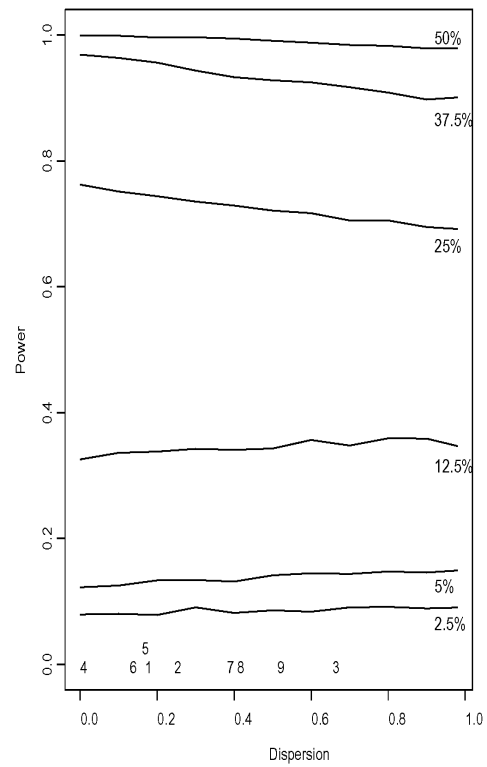


# Calculation of power

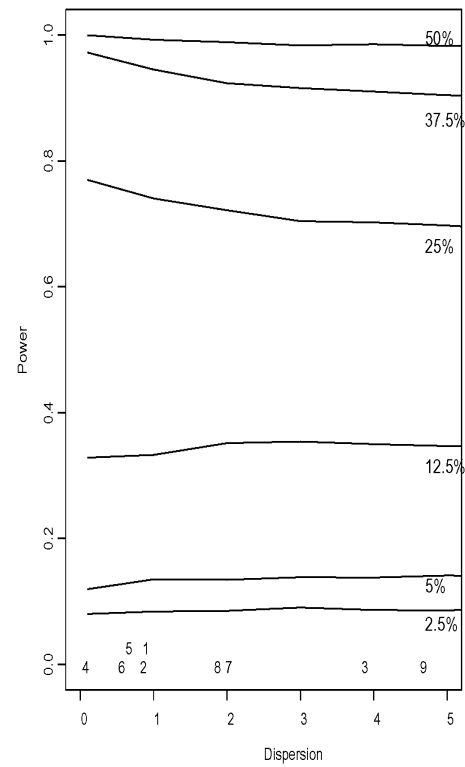
- ◆ Calculation of  $P(X = x_{\text{critical}})$
- ◆ Done by Monte Carlo methods:
  - Simulate  $X$  (total number of correct answers)
  - Count how often  $X$  is larger than or equal to the critical value
- ◆ Easy if software has built in functions:  
normal, beta, binomial. (e.g. Splus/R/SAS)

# Power, $n=12$ , $k=4$

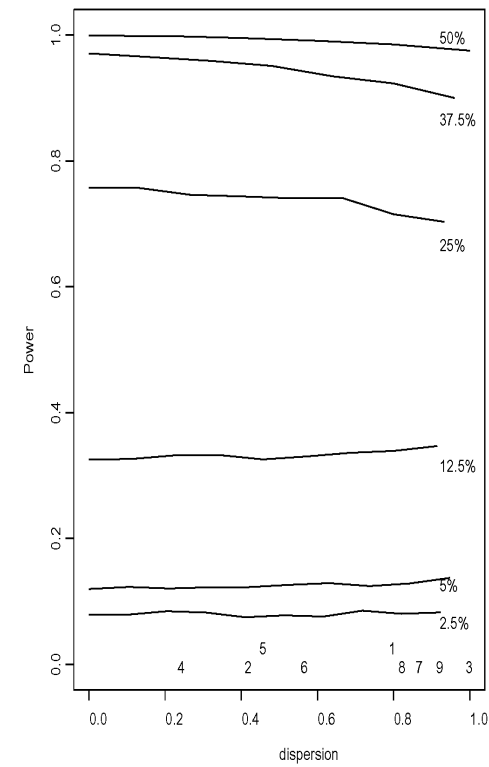
Mixture



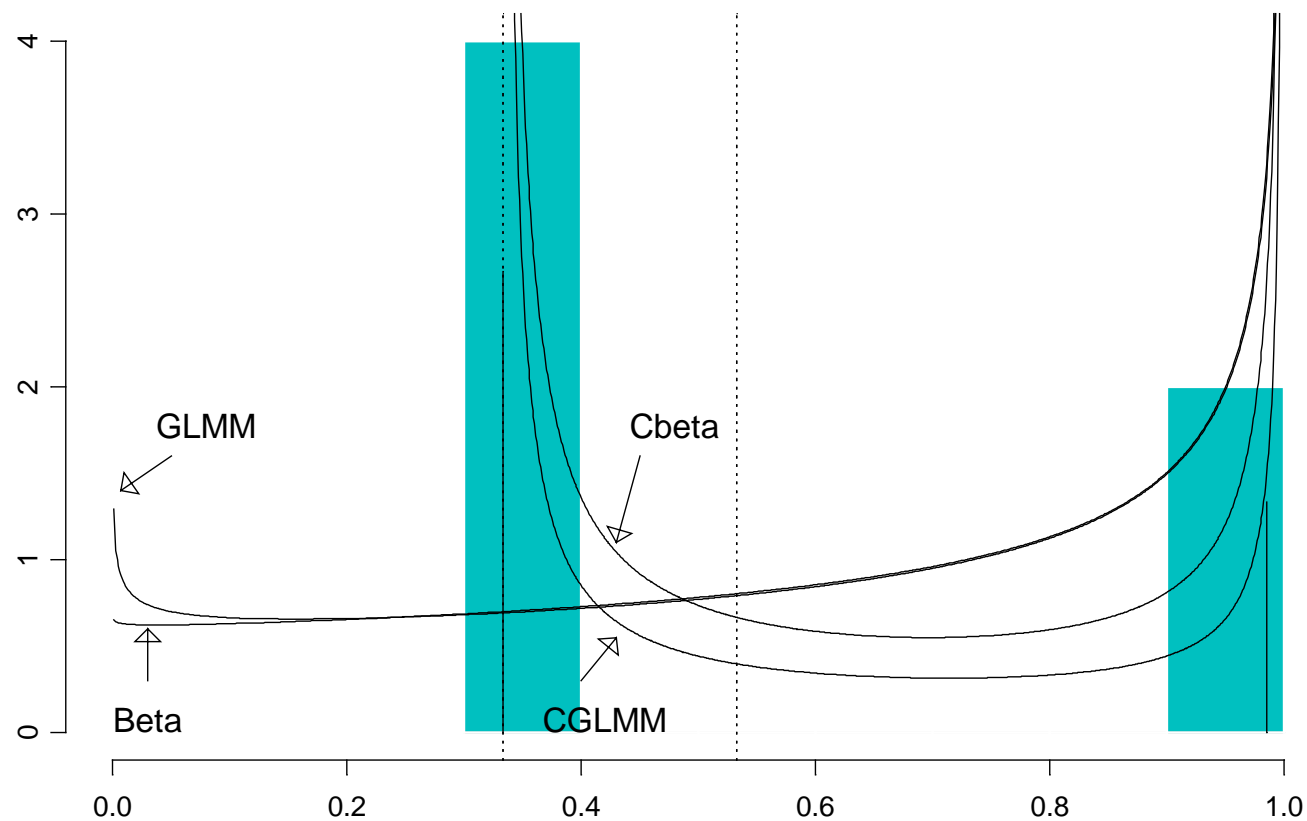
CGLMM



CBeta



# Extreme data: $n=6$ , $k=100$



# The common limit model

- ◆ All 3 models converge to the same limit/extreme situation:
- ◆ All individuals are either 100% discriminators or non-discriminators (guessers)

# The common limit model

For each assessor:  $X_i \sim \text{bin}(k, p_i)$

The individual probabilities are randomly distributed:

$$p_i = \begin{cases} \frac{1}{3} , & \text{with probability } 1 - \gamma \\ 1 , & \text{with probability } \gamma \end{cases}$$

# Monte Carlo for the common limit model

- ◆ Fix the effect size
- ◆ Simulate  $n$  binomial(1, ?)s
- ◆ For each outcome of 1 set  $x_i = k$
- ◆ For each outcome of 0 simulate a binomial( $k, 1/3$ )
- ◆ Count the number of times  $X$  becomes as large as the critical value.

# Limits of power for triangle test

Level 5%, 37.5% effect

n:	k=1	k=2	k=3	k=4	k=5
12	40%	70%	81%	90%	91%
.					
38	88%				
39	92%				
48	97%				

# Summary

- ◆ There is no big difference between the different “complicated models” to handle replications
- ◆ The loss of power by substituting assessors by replications is remarkable small
- ◆ Given the panel size, a few replications increase the power considerably
- ◆ Tables of limit power is given for some situations – a simple Monte Carlo method for other.

## Some additional insights

- ◆ Computationally, the mixture model is the easiest to handle:
  - The EM-algorithm easy to implement AND gives the option of "fuzzy clustering" of the individual assessors
- ◆ More powerful test than the independent binomial exist!